



# More effort with less pay: On information avoidance, optimistic beliefs, and performance<sup>☆</sup>

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## ABSTRACT

Neoclassical theory presumes that agents value instrumental information. In contrast, recent behavioral studies motivate and model information avoidance. We study preferences for and against instrumental information in a real-effort task varying information structures on performance pay. Our study offers three main results. First, we confirm that both, preferences for and against instrumental information, exist. Second, information avoiders outperform information receivers. This result holds independently of effects of self-selection. Third, our findings about information avoiders can be aligned with behavioral theories of optimistic belief design.

## 1. Introduction

Neoclassical theory assumes that agents are interested in instrumental information. When working, for example, agents would want to know the exact payment scheme such that they can adapt their performance accordingly, balancing costs and expected rewards.

Recently, this view on preferences for instrumental information has been challenged. Information avoidance seems to be a pervasive phenomenon in a variety of contexts. Two recent surveys (Golman et al., 2017; Hertwig and Engel, 2016), highlight a number of situations in which agents actively “decide not to know” and prefer to remain ignorant about instrumental information. Evidence stems from the health domain (Oster et al. (2013), Ganguly and Tasoff (2017), Serra-Garcia and Szech (2020), contexts relevant for one’s self image (Eil and Rao, 2011; Dana et al., 2007; Serra-Garcia and Szech, 2022; Chen and Heese, 2021) but also from financial domains such as portfolio investment decisions (Karlsson et al., 2009; Sicherman et al., 2016).

Recent models from behavioral theory on information avoidance can be divided into two groups. Either they assume that agents bias their expectations optimistically under coarse information structures, as in Brunnermeier and Parker (2005) or Gollier and Muermann (2010) or, they assume that agents have Bayesian beliefs. Examples for the latter case include (Bénabou and Tirole,

<sup>☆</sup> This paper was written before the tragic death of Nora Szech. Both of the other co-authors vividly remember the joyful and intellectually stimulating collaboration with her on this project. This paper is dedicated to Nora.

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2002) and Carrillo and Mariotti (2000). These models then rationalize information avoidance as a response to self-control problems induced by precise information.<sup>1</sup> Only under the former approach, uncertainty provides some leeway for unrealistic optimism.

The goal of this paper is three-fold. First, we explore whether preferences for information avoidance can also be found in a work-related context. To this purpose, we consider a real-effort task and vary how much information subjects have on performance pay. Second, we analyze how information avoidance affects performance. Our design allows us to contrast performance in self-selected (coarse) information structures versus performance in identical yet enforced information structures such that we can identify the causal role of (a lack of) information for performance. Third, we investigate the mechanism driving performance under information avoidance.

A number of applied theoretical papers have explored situations in which agents may want to avoid instrumental information, e.g., if anticipations influence utility. One prominent example is medical testing, which is often avoided by patients (compare Caplin and Leahy, 2004; Schweizer and Szech, 2018). Kőszegi (2003) considers a related problem in which doctors bias information delivered to patients, who accordingly infer that they are part of a cheap talk game. In all these approaches, optimal expectations balance the psychological benefits of deliberately coarse information with the material costs of making suboptimal choices. Data by Oster et al. (2013) on tests for the hereditary Huntington's Disease confirm that many people avoid instrumental information.<sup>2</sup> A large fraction of people at risk shy away from the genetic test, even though direct costs of testing are small and subsequent adjustments to test results are typically large. In addition, her data display that patients who avoid the medical test seem to do specifically well in life. The authors suppose that test avoiders are optimistic and tend to believe that chances of getting Huntington's are much smaller than the Bayesian correct expectation. Yet Oster et al. (2013) can only analyze self-selected information structures. Their findings may be driven by people with a more positive attitude towards life, self-selecting into avoiding the test. In our less drastic set-up, we are able to control for effects of self-selection.

In the real-effort context we study, agents know that there is a piece rate for the task that they perform and that this piece rate is either high (1 EUR) or low (0.1 EUR), with equal probability. Our results show that information about this piece rate is instrumental: The 1 EUR piece rate induces a significantly better performance than the 0.1 EUR piece rate. Nevertheless, we find that preferences for information avoidance exist, which is our first main result. When given the choice, about one third of agents opt against information about the piece rate. Second, we find that information avoiders perform better than information receivers. This effect holds even if information structures are enforced instead of self-chosen. Coarse information thus motivates agents better than precise information. Third, we elicit agents' performance at a fixed wage of 0.55 EUR. This corresponds to the expected wage of the uncertain payment scheme when there are no belief distortions. Under this certain piece rate, agents perform significantly worse than under the uncertain payment. This suggests that agents do not stick to the objective prior when uncertain about the piece rate.

Next, we try to calibrate optimistic belief distortions in our setting. Therefore, we implement another treatment. We choose a piece rate of 0.8 EUR because (a) the former treatments show that performance under an uncertain piece rate is at least as high as performance when the piece rate is known to be high, i.e., 1 EUR, and (b) in our ex-post questionnaire, several agents argue in favor of information avoidance in order to reduce feelings of pressure. Indeed, the phenomenon of choking under pressure has received considerable attention in the psychology literature since Baumeister's seminal (1984) paper. Choking has also been documented in a number of economic studies (see, for example, Dohmen, 2008; Ariely et al., 2009; Apesteguia and Palacios-Huerta, 2010). Therefore, a still high yet below 1 EUR piece rate could motivate agents best to perform well. We find that the 0.8 EUR treatment indeed leads to comparably good performance results as the uncertain piece rate scheme with an expected piece rate of 0.55 EUR. From a principal's view, the uncertain payment scheme thus dominates: with a substantially lower average piece rate, the same performance can be induced.

All in all, our results indicate that uncertainty can motivate agents very well. Indeed, tournament incentives which are widely used at the workplace might exploit this very mechanism. Likewise, the coarseness of payment contracts involving shares of the employing company may affect performance results—even in workers on lower levels of a company's hierarchy who are unlikely to influence a company's share-value through their individual behavior. If there is some room for optimism, it may well be the case that workers bias their beliefs in that direction. For a principal, it may thus be profitable to leave some elements in payment schemes undetermined. Likewise, employees may find it difficult to realize whether searching for a new job would be advantageous. It may be that chances for promotion or high bonus payments are often perceived too optimistically at the current employer.

Our paper is organized as follows. Section 2 describes the design of our experiment and formulates our key predictions, contrasting models which assume optimistic belief distortions with models that rationalize information avoidance as a strategic way to sustain sufficient motivation without distorting beliefs. Section 3 presents the results and argues that they, by and large, support the former approach. Based on this, Section 4 aims to pin down further the exact mechanism driving our results. In particular, we discuss other possible explanations for our findings, but ultimately argue that our data can best be rationalized by augmenting a model of optimal belief distortions by an additional channel in which agents' performance may suffer when wages are too high. Section 5 situates the paper in the literature and Section 6 concludes.

<sup>1</sup> Models of anticipatory utility in Caplin and Leahy (2001, 2004) also fall into this category as they do not allow for belief distortions.

<sup>2</sup> Huntington's is a genetically caused, severe disorder that breaks out around age 40, and leads to a drastic physical and mental decline.

## 2. Experimental design and hypotheses

We conduct four main treatments, FULL INFO, NO INFO, INFO CHOICE, and FIXED WAGE. The first three treatments are identical except that information about piece rates varies—exogenously or endogenously. The FIXED WAGE treatment allows us to separate between behavioral models on information avoidance assuming undistorted beliefs, such as Bénabou and Tirole (2002) and Carrillo and Mariotti (2000), and models assuming optimistic belief distortion à la (Brunnermeier and Parker, 2005). This treatment is therefore designed differently, as explained in detail below.

In FULL INFO, NO INFO, and INFO CHOICE, subjects know that they receive either a high (1 EUR) or a low (0.1 EUR) piece rate, with equal probability, for performance results in a tedious task.<sup>3</sup> Subjects have 60 min to enter lines of strings, containing numbers, upper case and lower case letters, *backwards* into the computer interface. For each correctly entered string, they receive a piece rate. Each string consists of 60 characters. For example, one of the strings used in the experiment looks as follows.

NXgCX7JHxYZj2cfBSd8JtkYp3LPcyDX8y8NNQhrzJfg22S2ACjC85EQ43B7L

Each task consists of one of these randomly generated strings, and all tasks are identical for all subjects.

We vary across treatments how much information subjects have about their piece rate while working on the task. In treatment FULL INFO, every subject gets to know their piece rate immediately, before working. In treatment INFO CHOICE, subjects can choose whether they want to know their piece rate immediately, or whether they only want to find out about it after they have finished working on the task. Thus subjects can avoid information while working. In treatment NO INFO, subjects are forced into information avoidance, but know that both piece rate realizations are equally likely. They thus know exactly as much as self-selected information avoiders in INFO CHOICE. At the end of the experiment, we ask subjects for some basic demographics about themselves.<sup>4</sup> In treatment INFO CHOICE, we also ask them in an open form about their reasons for choosing (not) to obtain information about their piece rate before working on the task.<sup>5</sup>

Treatment FULL INFO allows us to find out whether information about the piece rate is instrumental. This is the case if agents perform better under the 1 EUR piece rate than under the 0.1 EUR piece rate, which is exactly what we find. In line with recent behavioral theory, it may nevertheless be the case that preferences for information avoidance exist in this context.

**Prediction 1.** *Preferences for information avoidance exist, though information has instrumental value.*

We should note that here and throughout the paper, we refer to information being instrumental in the classical sense, i.e. as leading to an adjustment in performance when learning the specific wage realization. Of course, once we observe agents avoiding information, this reveals that not learning the precise wage has benefits for them that outweigh any performance adjustments. We want to explicitly distinguish this case from situations where information is “non-instrumental” because effort does not adjust in response to information. Indeed, we are interested in whether we find evidence for information in settings where effort *does* adjust once precise information cannot be avoided.

Models which attribute information avoidance to problems of self-control, as well as approaches assuming optimistic belief distortions, can rationalize this outcome, albeit via different channels.<sup>6</sup> We provide formal proofs for the derivations of all our predictions in the appendix. In the main text, we will briefly sketch out the theoretical framework used and highlight the key differences between the approaches, and contrast them with orthodox benchmarks.

We consider an agent who receives a piece rate  $w$  for each correctly solved task. She chooses effort  $e$ , which we assume to correspond directly to the number of tasks solved, and comes at an increasing and convex cost  $c(e)$ . The agent's payoff from working on the task is therefore  $w \cdot e$ . In what follows, we will assume risk neutrality.<sup>7</sup> Hence, for some wage  $w \in \{w_L, \bar{w}, w_H\}$ , where  $\bar{w} = \frac{1}{2}(w_L + w_H)$ , a standard expected utility optimizer maximizes

$$w \cdot e - c(e).$$

We assume that an interior solution exists. This leads to the according optimal effort levels  $e_L, \bar{e}$ , and  $e_H$ . These are our orthodox benchmarks. We will now compare the predictions to two classes of models that differ in their rationale for why we might see information avoidance in our setting.

According to the first modeling approach, the agent may face a problem of temptation when she finds out that the wage is low (e.g. Bénabou and Tirole, 2002 or Carrillo and Mariotti, 2000). From an ex-ante perspective, she would like to exert the optimal effort level  $e_L$ . Yet, once she learns that the piece rate is low, she will choose an effort level of  $\hat{e}_L < e_L$  due to loss of motivation.<sup>8</sup> If

<sup>3</sup> At the time of the experiment, 1 EUR  $\approx$  1.37 USD.

<sup>4</sup> A full list of characteristics can be found in Table A.1 in the appendix, together with the corresponding summary statistics by treatment.

<sup>5</sup> The exact wording of the question was “Please explain why you decided to look at the colored cardboard at the beginning of the experiment, or respectively, why not”. The colored cardboard was used to implement information revelation, see below.

<sup>6</sup> See also Bénabou (2015) for a comparison of the two approaches. For other applied studies documenting preferences for information avoidance, see the overviews by Golman et al. (2017) and Hertwig and Engel (2016).

<sup>7</sup> In Section 4, we provide an argument why relaxing this assumption does not affect our predictions.

<sup>8</sup> We assume that this decrease in motivation only happens for wages that are sufficiently low, i.e. below  $\bar{w}$  in our case. While we leave unspecified at what wage exactly the self-control problem of exerting too little effort (from an ex-ante perspective) kicks in, we make this assumption because it best captures the temptation of slacking off. In addition, having to work for the low wage is the result of a random event, which makes the comparison to the forgone high wage very explicit and the motivation problem more severe.

she anticipates this effect, she may prefer to avoid information and exert an effort of  $\bar{e}$  instead. Information avoidance thus serves as a commitment device. If the expected piece rate  $\bar{w}$  is sufficiently attractive, the agent may decide to stick to this expectation instead of receiving full information and running the risk that motivation decreases severely. We formalize this argument in [Appendix A](#).

In models with belief distortions such as [Brunnermeier and Parker \(2005\)](#) or [Gollier and Muermann \(2010\)](#), an agent derives part of her utility from anticipation, that is, utility based on a (possibly distorted) belief about realizations of future utility. Applied to our real-effort setting, these models assume that agents optimistically distort their beliefs about their piece rate under uncertainty. Agents trade off the benefits from biased beliefs about the payments against the costs of ex-post suboptimal effort choices. This distorts effort choices as follows. When choosing to avoid information, agents exert effort  $\bar{e} > \bar{e}$  because they choose the optimal effort level based on optimistically distorted beliefs about the likelihood that the high wage realizes. Hence, if agents place sufficiently large weight on their belief utility, they will avoid precise information, and prefer to leave some room for optimistic belief distortion. This may happen consciously, or sub-consciously. Again, the appendix contains the formal details.

The crucial difference between the two behavioral modeling approaches for our context is whether agents optimistically distort their expectations and perform accordingly, or whether their performance suggests that they maintain the undistorted objective beliefs instead. Before outlining how our design can distinguish between the two, we note that both approaches support that coarse information can boost performance compared to precise information. This leads us to our second main hypothesis.

**Prediction 2.** *Coarse information causally fosters performance results compared to precise information, even though expected piece rates (based on the objective prior) are identical in both settings.*

Both classes of models that we discuss can explain a preference for information avoidance, and an according increase in performance compared to the orthodox benchmark under uncertainty. The difference in predictions comes when analyzing performance results under coarse information in more detail. According to approaches that focus on the self-control problem, agents' performance when the precise wage is unknown is based on the undistorted expectation. In contrast, in models of optimistic belief distortion, agents hold behaviorally biased expectations and perform accordingly.

We implement the **FIXED WAGE** treatment in order to test for these different predictions. In **FIXED WAGE**, agents know that they earn a piece rate of 0.55 EUR for sure. This is the expected wage from the **NO INFO** treatment. Without belief distortions, agents should have this expectation in mind while working in **NO INFO**. Performance results should then be comparable to those from **FIXED WAGE**. If there was a difference in performance, it should be that **FIXED WAGE** outperforms **NO INFO** as agents typically tend to be risk-averse.

In contrast, when agents hold too optimistic beliefs regarding the future payments under coarse information, they could thus perform significantly better in **NO INFO** than in **FIXED WAGE**. If so, from a principal's point of view, keeping the piece rate uncertain could causally foster performance compared to a certain payment of same expected value.

**Prediction 3.** *If agents behaviorally distort their beliefs under uncertainty, the number of correctly solved tasks in **NO INFO** can be higher than in **FIXED WAGE**. If the opposite holds true, the data support approaches that explain information avoidance as a way to overcome self-control problems to maintain sufficient motivation when wages are low.*

In contrast to field studies, e.g., from the literature on medical testing, our experimental design can further shed light on whether there is a *causal* effect of coarse versus precise information on behavior. This is possible because part of our subjects are forced into specific information structures, as is the case in **FULL INFO** as well as in **NO INFO**. Our set-up therefore allows us to isolate the causal role of the information structure on work performance. To measure potential additional effects of self-selection, we compare performance from **NO INFO** and **FULL INFO** to performance in the corresponding, self-selected sub-groups from treatment **INFO CHOICE**.

In the rest of this section, we describe the implementation of the treatments in detail. We present the results from our study in [Section 3](#).

**Implementing the different information structures across treatments.** The randomization of subjects into the treatments **FULL INFO**, **NO INFO**, **INFO CHOICE** and **FIXED WAGE** was done at the session level. For the three treatments **FULL INFO**, **NO INFO** and **INFO CHOICE** the detailed procedures employed to determine a subject's individual wage were as follows. When entering the lab, each subject is randomly allocated a red or a black chip by one of the experimenters. Half of the chips are black, the other half red. Each subject is then told to take a seat at a computer terminal where the screen shows a square with the color corresponding to the color of his or her chip.<sup>9</sup> Subjects know that depending on the color (red or black), they can either earn 0.1 EUR or 1 EUR for each correctly entered string, with equal probability.

In order to determine which color corresponds to a high piece rate and which to a low one, we use the following procedure. We prepared two pieces of cardboard which look identical from the outside when folded, but inside either show a red or a black square. After showing the cardboard pieces (from outside and inside) to all participants, they are folded, secured with paper clips, placed into a small bag and shuffled. Another experimenter then draws one of the two folded cardboard pieces. The color of the drawn piece determines which color is associated with the high wage for this session.

In treatment **FULL INFO**, the cardboard is unfolded and the color is revealed to all subjects immediately. Thus subjects know whether they are going to receive the high piece rate, or the low one.

In treatment **NO INFO**, the folded cardboard is placed onto a white board at the front of the room where it remains for the whole duration of the experiment and is revealed to all participants once the allowed time for the task (60 min) is up. Hence subjects do

<sup>9</sup> The procedure makes sure that subjects are aware that there are chips of two colors and that the procedure is entirely random.

not know whether they earn the high or the low piece rate while working on the task.

In treatment **INFO CHOICE**, subjects are asked on their computer screen whether they would like to receive the information about the color now, or wait until the end of the experiment. After clicking the button corresponding to their choice, another screen appears which states the subject's decision. After all subjects have made their choice, the experimenter walks through the lab and privately reveals the color inside the cardboard to those subjects who decided to see it. As in the **NO INFO** treatment, the folded cardboard is then placed onto a white board and revealed afterwards. Thus, in this treatment, subjects choose whether they want to know their piece rate beforehand or not. Then the real effort task starts.

After each string that subjects enter, they learn whether they entered it correctly. They can then click on a button to continue. Subjects are informed about the time that remains. Subjects are not allowed to use any electronic devices, but are each given a copy of a well-known German weekly, called *DER SPIEGEL*. This magazine has a weekly circulation of more than one million. It contains all sorts of articles, from investigative journalism including reports on politics to articles about scientific discoveries and information on cultural events and sports. Subjects are explicitly told that they can make use of the magazine "...whenever, during the experiment, [they] would like to take a break or pass time". Thus, no subject has to feel obliged to work on the task if he or she prefers to spend their time otherwise.

In the **FIXED WAGE** treatment, there is no need for any randomization at the beginning and subjects immediately start working on the task after reading the instructions. We also inform subjects that other participants could earn either 0.1 EUR or 1 EUR for the same task in previous experiments, with equal probability. Giving this additional context to participants ensures that the **FIXED WAGE** treatment induces similar social comparisons as in the other treatments. In particular, this ensures that in all our treatments participants are aware of the highest and the lowest possible wage, i.e., 0.1 EUR and 1 EUR, that other experimental participants could have possibly received.

In total, our sample consists of 238 subjects. All treatments were run at the WZB-TU Laboratory in Berlin between November 2013 and April 2014. There were no restrictions imposed on the invited participants regarding gender, subject of study, or previous experience with experiments. We used z-tree (Fischbacher, 2007) as the experimental software and ORSEE (Greiner, 2015) to recruit subjects. Participants received a show-up fee of 5 EUR and average earnings over all treatments amounted to 14.29 EUR. Each session lasted 80 to 90 min.<sup>10</sup>

### 3. Results

In treatment **FULL INFO**, subjects know their piece rate realization before starting to work on the task. Our data show that subjects perform significantly better under the high than under the low piece rate. Subjects working for the low piece rate of 0.1 EUR solve 20.67 tasks on average, whereas subjects working for the high piece rate of 1 EUR solve 26.21 tasks correctly. This difference is significant ( $p = 0.043$ ),<sup>11</sup> and confirms that information about the piece rate is instrumental as hypothesized.<sup>12</sup> The result furthermore shows that the combination of the task we employed and the difference between the low and the high wage is suitable to uncover meaningful responses to incentives, despite the relatively large number of solved tasks in the low wage condition.

Having established that information is instrumental, we can turn to Prediction 1. In support of all theories discussed which predict information avoidance, 30 out of 95 subjects (31.6%) prefer not to know their piece rate while working. These subjects thus decide to forgo information of instrumental value as shown in treatment **FULL INFO**.<sup>13</sup> The decision to avoid information does not correlate with any of the demographic characteristics we elicited, as we show in Table A.2 in Appendix A. We are also confident that the decision to avoid information is not driven by confusion, inattention or noise as our design took utmost care that participants had to make an active choice whether or not to avoid information and that this choice was presented to them in a clear and unambiguous manner.<sup>14</sup> In addition, as we highlight in more detail below, when we ask participants in the post-experimental questionnaire for their reasons for avoiding information, their answers are of high quality, suggesting that at least the vast majority of participants made a well-informed decision. Our data thus confirm Prediction 1.

Next, we turn to Prediction 2. The findings from **INFO CHOICE** support Prediction 2. Information avoiders solve 30 tasks correctly on average, while information receivers solve only 21.31 tasks correctly ( $p = 0.0002$ ). Controlling for potential effects of self-selection, a similar picture emerges: **NO INFO** leads to significantly better performance results than **FULL INFO** (23.44 versus 28.02,  $p = 0.0274$ ). We thus confirm Prediction 2. Information avoidance increases performance. Those information receivers who learnt that their piece rate is 1 EUR solve 25.53 tasks on average. When we test whether this is statistically different from the 30 tasks solved by information avoiders, the test misses statistical significance ( $p = 0.0573$ ). From a principal's point of view, giving agents a choice to

<sup>10</sup> We also ran two sessions without the weekly magazine. We report the findings from these sessions as well. Overall, results are quite similar to those in our main treatments, though effects are slightly less pronounced.

<sup>11</sup> Unless indicated otherwise, all p-values are calculated using a Wilcoxon Rank Sum test.

<sup>12</sup> The other possible measure of performance would be the number of attempted tasks. Arguably, our task is prone to errors and for some subjects this measure might more correctly reflect the actual effort put in. Others might choose a more risky strategy and tolerate more errors. The two measures are highly correlated ( $\rho = 0.8424$ ,  $p = 0.0000$ ) and the results are very similar.

<sup>13</sup> As a robustness check, we analyze the two sessions that were identical to **INFO CHOICE** except that subjects did not have access to the magazine. There, 15 out of 44 subjects (34.1%) decided not to acquire information, an effect of almost identical magnitude.

<sup>14</sup> As highlighted by Exley and Kessler (2023), requiring participants to make an active decision to avoid information can substantially reduce the share of participants avoiding information. Their experiments show that in the moral wiggle room paradigm of Dana et al. (2007), a substantial amount of information avoidance can be attributed to factors like confusion or inattention. They find that in their setting, the degree of information avoidance is substantially reduced if an active choice to avoid information is required and the decision is separated from other decisions in the experiment.



**Table 1**  
Mean and median performance across treatments.

| Piece rate | FULL INFO        |                       | INFO CHOICE      |                       | NO INFO         |                       | FIXED WAGE      |                       |
|------------|------------------|-----------------------|------------------|-----------------------|-----------------|-----------------------|-----------------|-----------------------|
|            | Mean<br>(s.d.)   | Median                | Mean<br>(s.d.)   | Median                | Mean<br>(s.d.)  | Median                | Mean<br>(s.d.)  | Median                |
| 0.1        | 20.67<br>(10.49) | 22<br><br><i>N=24</i> | 17.69<br>(11.37) | 17<br><br><i>N=35</i> |                 |                       |                 |                       |
| 0.55       |                  |                       |                  |                       |                 |                       | 24.66<br>(9.58) | 24<br><br><i>N=47</i> |
| 1          | 26.21<br>(8.75)  | 26<br><br><i>N=24</i> | 25.53<br>(9.86)  | 23<br><br><i>N=30</i> |                 |                       |                 |                       |
| unknown    |                  |                       | 30<br>(9.35)     | 29<br><br><i>N=30</i> | 28.02<br>(8.41) | 28<br><br><i>N=48</i> |                 |                       |
| aggregate  | 23.44<br>(9.17)  | 25<br><br><i>N=48</i> | 24.05<br>(11.44) | 23<br><br><i>N=95</i> | 28.02<br>(8.41) | 28<br><br><i>N=48</i> | 24.66<br>(9.58) | 24<br><br><i>N=47</i> |

avoid information thus does not harm performance. Instead, information avoiders perform especially well.<sup>15</sup>

While coarse information significantly enhances performance, effects of self-selection are not significant in our data. We test for this by comparing for the known wages ( $w = 0.1$  and  $w = 1$ ) whether performance is significantly different between treatment INFO CHOICE and treatment FULL INFO and find no difference (20.67 versus 17.69,  $p = 0.2138$ , and 26.21 versus 25.53,  $p = 0.3975$ ). The same holds for the case where the wage is unknown when we compare performance in INFO CHOICE to performance in NO INFO (30 versus 28.02,  $p = 0.3710$ ). In addition, in Table A.3 the appendix, we confirm that the same result holds when using OLS regressions instead. Therefore, in the following, we sometimes also consider the sample where we pool self-selected and enforced information together in order to facilitate focusing on the main effects.

So far, the data are in line with both approaches that rationalize information avoidance through optimistic belief distortion as well as models which maintain undistorted beliefs. The difference between the two model approaches is reflected in Prediction 3. According to the latter approach, under coarse information, subjects should perform worse or equally well as when they work for the expected wage paid out for sure. To the contrary, if we allow for optimistic belief distortions, subjects may perform better under uncertainty than under the secure piece rate matching the undistorted expectation. The latter is what we find.

In FIXED WAGE, subjects know that they work for a piece rate of 0.55 EUR. They also know that other subjects worked for either 1 EUR or 0.10 EUR with equal probability, in order to keep the context as comparable as possible. We find that subjects solve on average 24.66 tasks correctly in this treatment. This is less than the average of 28.78 tasks if the wage is unknown and 0.55 in expectation. The difference is highly significant ( $p = 0.0015$ ). With regard to our two classes of models, this increase in performance under uncertainty is only compatible with approaches of optimistic belief distortion such as Brunnermeier and Parker (2005). Under models without belief distortions, such as Bénabou and Tirole (2002), we would have to see the opposite effect if agents tend to be risk-averse. Even with risk-neutral agents, there is no reason why agents in NO INFO should outperform those in FIXED WAGE. Yet we see that this is the case. For a principal, implementing a situation of coarse information fosters performance compared to providing the expected piece rate with certainty.

Table 1 presents a detailed overview of the performance in the different treatments and for the different piece rates. Table 2 displays the performance results for the different information structures subjects had while working on the task, pooling the data of self-chosen and enforced information structures as there are no significant differences between the two at any conventional level. We find that the result from above regarding performance for the known wages 0.1 EUR and 1 EUR carries over. Subjects perform significantly better at the higher piece rate (25.83 vs. 18.90 correct tasks on average,  $p = 0.0003$ ).<sup>16</sup> In Table A.4 in the appendix, we present the results from OLS regressions, analogous to Table 2, showing that our findings are robust to the inclusion of individual characteristics as additional control variables.

Fig. 1 plots the empirical distribution functions for the four cases of information structure. We see that under the uncertain piece rate, performance results are best. All in all, we obtain a strong indication that uncertainty fosters performance compared to paying the expected piece rate with certainty. Visually, the distribution for the case in which the piece rate is unknown almost stochastically dominates all other distributions. Paralleling our finding in the INFO CHOICE treatment, the  $p$ -value of the test for a difference between the distribution for the 1 EUR piece rate and the unknown piece rate is given by  $p = 0.0645$ .

<sup>15</sup> Table A.3 in the appendix confirms these results when we use OLS regressions instead, where we also control for observable individual characteristics.

<sup>16</sup> It appears that most subjects did not take longer breaks during the one hour working time. Even under the low piece rate of 0.1 EUR, only four subjects spend more than ten minutes on any task. Looking at the other piece rates (including unknown), no subject paused for a longer period of time; they always move on to the next task within six minutes.

**Table 2**  
Pairwise comparisons of performance results across piece rates.

| Piece rate | POOLED           |        | p-value of pairwise test |        |        |
|------------|------------------|--------|--------------------------|--------|--------|
|            | Mean<br>(s.d.)   | Median | 0.1                      | 0.55   | 1      |
| 0.1        | 18.90<br>(11.03) | 19     |                          |        |        |
| 0.55       | 24.66<br>(9.58)  | 24     | 0.0091                   |        |        |
| 1          | 25.83<br>(8.56)  | 25     | 0.0003                   | 0.2073 |        |
| unknown    | 28.78<br>(8.78)  | 28     | 0.0000                   | 0.0015 | 0.0645 |

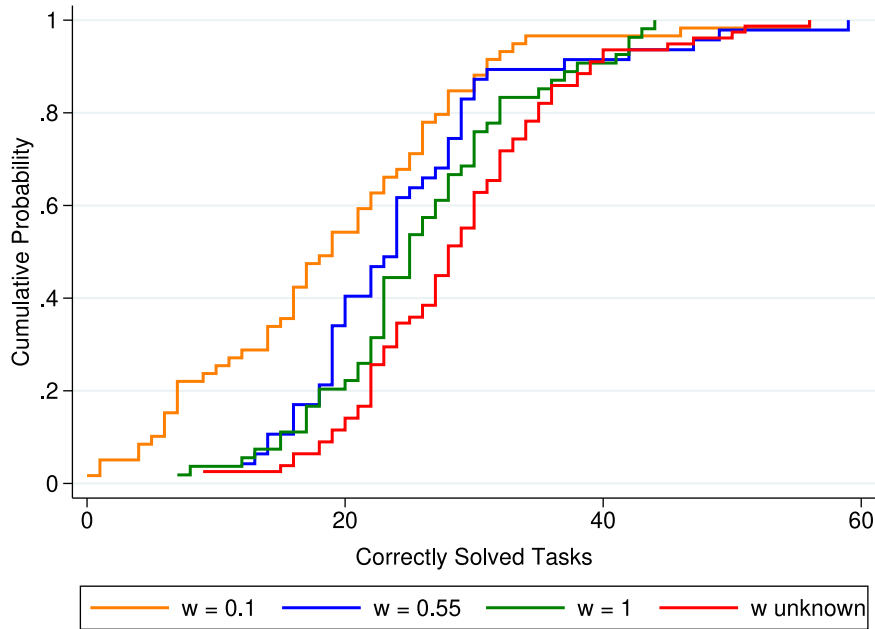


Fig. 1. CDFs of performance, by wage.

From a principal's perspective, the data show that giving people the choice to avoid information does not harm aggregate performance compared to a full information policy. Subjects who self-select into information avoidance perform extremely well. This result is in line with findings from [Oster et al. \(2013\)](#). Moreover, enforcing information avoidance significantly increases performance results. This effect cannot be explained by theories that maintain correct expectations. Instead, agents may optimistically bias their beliefs, as we will discuss in detail below. If agents bias their beliefs, the question emerges, of course, which optimistic beliefs agents may have in mind. This we investigate in more detail in the following section.

#### 4. Discussion

In this section, we address four points. First, we explore in more detail the potential role of optimistic belief distortions in the spirit of [Brunnermeier and Parker \(2005, BP henceforth\)](#), because our results suggest that agents bias their beliefs to values different from the undistorted expectation. Second, we report on a follow-up experiment we ran in order to roughly calibrate the magnitude of belief distortions. Third, we discuss possible alternative explanations for our findings. Fourth, we share some thoughts regarding the design of incentives.

##### 4.1. Combining belief design and choking

As discussed above, a model of behaviorally distorted beliefs can predict both a preference for information avoidance as well as an especially high performance when the wage is uncertain. Yet our performance results under uncertainty are nevertheless very pronounced. Specifically, under uncertainty agents perform at least as well than under the best possible realization of the uncertain

payment scheme. Under the BP approach, in order to rationalize such behavior, we would have to assume that agents are convinced that they receive the high wage for sure. Even though we did not elicit beliefs directly, this seems unlikely to be the case. Therefore, the question emerges whether within the BP framework there exists a belief which is optimistically biased, but not to the fullest possible extent. In the following, we will see that this kind of optimistic belief distortion to an expected piece rate below 1 EUR can serve as the best motivation for agents if they are prone to choke under pressure.

In the post-experimental questionnaire we conducted, our subjects provide different reasons for why they avoided information. Two arguments for information avoidance come up frequently. About a quarter of subjects (7 out of 30) argue that they wanted to avoid knowing that the piece rate realized low instead of high in order to feel more motivated. Yet, a similar share of participants (8 out of 30) argue that they intended to avoid feeling stressed and to prevent choking under pressure.<sup>17</sup> Overall, it is reassuring that the vast majority of participants seem to have thought through their decision to avoid information about their wage. At most three subjects give reasons that suggest that they were largely indifferent whether to avoid information or not and there is not a single subject whose answer suggests that they were confused about their choice or did not know the consequences.<sup>18</sup>

The fact that the idea of choking under pressure is mentioned prominently in the subjects' answers in a free-form questionnaire, i.e., without us pointing subjects specifically towards this issue, motivates us to further explore a model which combines belief design à la Brunnermeier and Parker (2005) with choking under pressure. To be clear, when doing so we simply aim to demonstrate that choking under pressure can, in principle, explain our results, unlike a number of other possible explanations which we consider in Section 4.3. We do not claim that choking under pressure is the only possible explanation, but it is one that we feel worthy of exploring a bit deeper in what follows.

Our approach to modeling choking — which we set out formally in the appendix — assumes that for a wage  $w$ , an agent's "output", i.e. the number of correctly solved tasks, depends on both the effort  $e$  she puts in and her productivity  $f(w)$ , which, in turn, depends on the wage she is facing. Importantly, while productivity may be constant or increasing for low or medium wages, once the stakes get too high, productivity declines. For a known wage, this implies that performance on the task will be "hump-shaped" in the wage that the agent receives. Hence, while for low wages an increase in the piece rate unambiguously increases effort and thereby output (which is, for example, what we see when comparing performance at 0.1 EUR to 0.55 EUR), very high wages can decrease output again (e.g., comparing piece rates of 0.8 and 1 EUR). As agents can actively distort their beliefs under uncertainty, an uncertain wage works very differently from a fixed wage. In particular, it allows agents on the one hand to make themselves more optimistic than the objective probability which increases effort and performance. On the other hand, they would not want to distort their beliefs too far as this would lead to potential choking and thus to a performance that may still be good, yet below the global optimum. Output at the optimal belief, say 0.8 EUR, could thus be larger than under the certain wage of 1 EUR.

We also note that for our results we do not need to assume that all agents choke under pressure. The same conclusions hold if we were to model a mixed population of chokers and standard agents whose productivity is non-decreasing over the range of wages we consider. This argument highlights an important role for heterogeneity in this setting: With heterogeneous agents, coarse information structures can outperform any kind of fixed wage scheme. The intuition behind this is that agents who work best if incentives are high have room to believe payments will be large. At the same time, agents afraid of choking under pressure have enough room for sticking to moderate expectations, thereby avoiding situations of choking.

We are fully aware that the stakes in our experiment are somewhat lower than in experiments which find evidence for choking under pressure, i.e. Ariely et al. (2009). Nevertheless, subjects who get paid 1 EUR per correctly solved task earn on average more than 25 EUR (plus a show-up fee of 5 EUR) for a session lasting 80 to 90 min, which corresponds to an hourly rate that is substantially higher than what subjects could typically expect to earn in these experiments in this lab at the time the experiments were run. Hence, given that our subjects have experienced previous experiments, it is conceivable that they consider a piece rate of 1 EUR somewhat special or at least unusual which might increase the pressure to do well for some subjects.<sup>19</sup>

#### 4.2. Additional treatment

We thus design an additional treatment which allows us to shed some more light on the role of optimal belief design in our setting, and how it interacts with potential choking. Specifically, we run a version of the treatment **FIXED WAGE** but change the wage agents receive from 0.55 EUR to 0.8 EUR. Otherwise, the treatment remains exactly the same, including the information that subjects in previous experiments worked for wages of 0.1 EUR and 1 EUR. The choice of 0.8 EUR corresponds to an expected wage which could result from an optimistic belief distortions under uncertainty. At the same time, it is sufficiently different from 1 EUR. This might reduce potential effects of choking if they exist. We run this treatment in November 2016 with 47 new subjects, at the same location and from the same subject pool as before.

<sup>17</sup> For the first group, a typical statement reads "I wanted to keep my motivation up with the hope of receiving the higher amount", whereas those potentially choking under the pressure of a high wage gave answers like "I didn't look at the card to take the pressure off myself. That would have made my work even more incorrect". The answers of all thirty participants who avoided information can be found in Online Appendix B.

<sup>18</sup> Among those who decided to obtain information, the majority of subjects (49 out of 65) state that they wanted to know their piece rate in order to adjust their effort accordingly. These subjects thus acknowledge the instrumental role of information for them. A typical statement was, for example, "I wanted to know whether putting in effort would be worthwhile, if the wage would have been low, I would not have bothered".

<sup>19</sup> In addition, our personal experience from doing the task ourselves and the feedback from colleagues with whom we tested the task when designing the experiment suggests that it is a task that requires substantial concentration. Also, the feature that if one of the 60 characters is entered wrong, the whole task is deemed incorrect and cannot be corrected adds (at least from our own perception) substantial pressure.



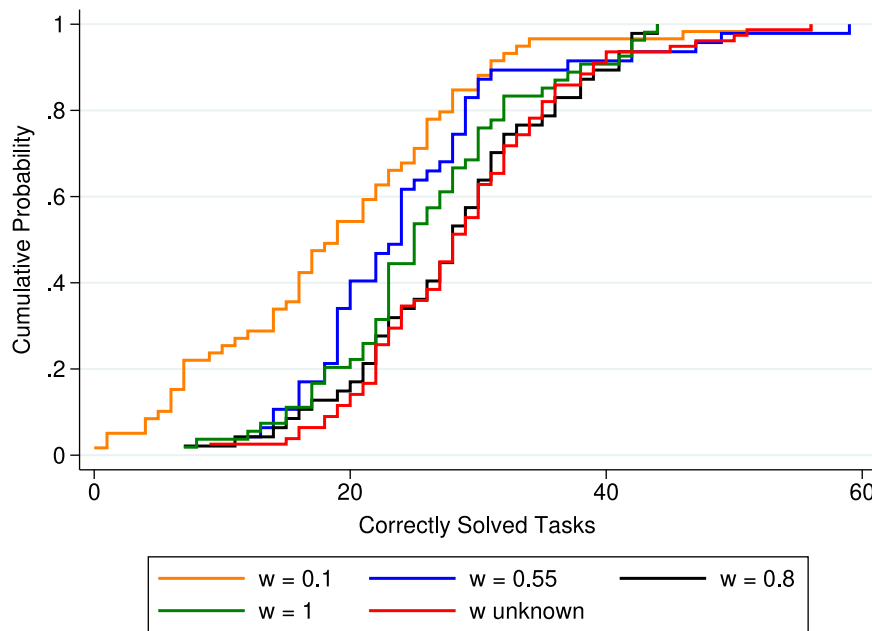


Fig. 2. CDFs of performance, by wage, including additional treatment.

We find that on average, participants solve 27.87 tasks correctly (median 28, s.d. 8.57). This is significantly higher than performance at 0.1 EUR and at 0.55 EUR, but we cannot reject the hypothesis that performance at 0.8 EUR is the same as under 1 EUR ( $p = 0.1953$ ). Fig. 2 plots the distribution functions of performance as in Fig. 1, enriched by the distribution from the new treatment. This new distribution is virtually indistinguishable from the distribution of performance when the wage is unknown. Thus it seems that the calibration at 0.8 EUR is not too far from the optimal uniform wage which a principal could set in order to mimic the performance results under uncertainty. There is, however, a natural limitation of any fixed wage scheme if different agents want to distort their beliefs differently, according to their individual needs. An uncertain wage scheme can leave room for agents to do so.

At this point, we should emphasize that while the data provide some evidence that choking under pressure might be at play in our experiment, the evidence for this is not fully conclusive on this issue. We have argued above how being prone to choking might interact with optimal belief design and in doing so could explain our somewhat puzzling finding that agents perform so well under uncertainty and find some support for this mechanism also in the subjects' statements. Nevertheless, we merely argue that choking under pressure seems a possible explanation, one that, to us, seems more plausible than others which we discuss in the following.

#### 4.3. Alternative explanations

In the following, we discuss other attempts to explain our empirical results. Specifically, we discuss risk preferences, satisficing, and regret aversion.

**Risk Preferences.** So far, we have considered risk neutral agents, both when deriving the theoretical predictions as well as when analyzing the results. A natural question to explore is whether risk aversion or risk lovingness could explain the empirical findings. The first, and most straightforward case to consider is the decision to avoid information. It is clear that a decision maker who behaves according to orthodox economic models would never avoid information in our setting, because information is always valuable to adjust effort. This in itself demonstrates the need for behavioral models to explain our data.

It is, however, less obvious that, conditional on having explained information avoidance, our predictions regarding performance that allowed us to distinguish between BP and BT still hold for agents who are not assumed to be risk neutral. We first note that risk seeking behavior could — in theory — explain a higher performance under uncertainty than under the fixed payment of 0.55 EUR. If agents have an explicit preference for uncertainty in their payments, they exert high effort to increase the spread in final payoffs. There is, however, little evidence for risk seeking behavior (see, e.g. Holt and Laury, 2002), either in the lab, nor in the field, which makes risk lovingness an unlikely — and arguably unsatisfactory — candidate to explain our results.

The converse argument implies that risk aversion should work against our findings. As the number of solved tasks increases, the variance in final payoffs increases, making exerting high effort unattractive. Importantly, this result continues to hold if we also consider the effects of higher order derivatives of the utility function. In particular, one could think that there might be cases where the agent is “prudent” (i.e.  $u'''(x) > 0$ ) and therefore feels the need to counterbalance increased risk with higher expected payoff, similar to a precautionary savings motive. Yet as long as optimal effort is concave in a fixed wage  $w$  (which is both theoretically

sensible to assume as well as clearly borne out by our data), the effect of prudence cannot outweigh the negative effect of risk aversion on optimal effort provision. We also note that for a risk-averse agent who engages in belief distortions, it is still the case that he would exert more effort under uncertainty than under the certain expected wage, provided the distortions are sufficiently large.

*Satisficing.* As argued by [Simon \(1955\)](#), agents may be satisficers instead of utility maximizers. That is, once payoffs reach some critical level, agents do not find it worthwhile to exert more effort. Coupled with optimally designed beliefs, satisficing could be an alternative explanation to choking: Agents may not exert as much effort when working for a wage of 1 EUR once they are “satisfied” with the number of tasks they have completed.

Two reasons speak more to the direction of choking, though. First, several subjects mention a fear of choking in the questionnaire, while no subject mentions ideas of satisficing. Second, there is no evidence for satisficing in the data either. For wages larger than 0.1 EUR, subjects never seem to stop working before the allotted time of 60 min is up. (In contrast, for a wage of 0.1 EUR, some subjects stop working.) Stopping once one has earned enough, however, would be a behavior expected from a satisficer.

*Regret Aversion.* Another possibility would be that agents feel regret when they work under wage uncertainty. This could lead them to perform extremely well in order to reduce regret should the low wage be revealed. The problem with this argument, and related concepts such as disappointment or loss aversion, is that working hard to avoid regret leads to a high spread of possible payments before the lottery is resolved. This is unlikely to be an attractive feature for agents. In addition, neither regret nor loss aversion can offer an explanation for the observed information avoidance. The model by [Gollier and Muermann \(2010\)](#) combines disappointment aversion and belief distortions, but, as the authors show, the former works as an opposing force to optimistic beliefs. Their model therefore predicts that agents have a preference for early resolution of uncertainty, contrary to what we find.

#### 4.4. Some thoughts on optimal payment schemes and information structures

Our results have implications for principals who want to design payment schemes for their agents. The high performance when the wage is uncertain suggests that such wage schedules are profitable for principals. They either increase output while keeping the expected wage constant (uncertain wage vs. 0.55 EUR), or keep performance at a similar level while lowering the payment to workers (uncertain wage vs. 0.8 EUR). This, however, begs the question whether a principal could do even better than using any of the payment schemes we investigated. We leave an in-depth investigation of this topic for future research but, nevertheless, offer some speculative thoughts.

Given that we find clear evidence for active belief design, it remains an intriguing possibility whether one could lower the expected wage of the lottery by making the low wage more likely than the high wage, while still obtaining similar levels of performance. Especially agents for whom anticipatory feelings are important, might respond by distorting beliefs to similar levels, thereby reducing effort only a little, but at the same time allowing the principal to save money (in expectation).

Another promising area to explore would be whether partial information release has positive effects on performance. For the context of medical testing, in which agents typically shy away from the test in fear of ending up in a situation of no hope, [Schweizer and Szech \(2018\)](#) explore possibilities of partial revelation. They find that for the medical context such as testing for Huntington’s, tests that perfectly reveal the good, but shroud the bad state may be a promising alternative for the vast majority of agents who do not dare to take the perfectly revelatory test, but have an interest in better information. This is even the case under undistorted beliefs. Similar types of information structures could be explored in the work context.

## 5. Related literature

We find that about one third of our subjects prefer to avoid instrumental information. For various but arguably very different contexts, preferences for information avoidance have been documented by [Golman et al. \(2017\)](#) and [Hertwig and Engel \(2016\)](#). The authors report preferences for information avoidance when it comes to self-image, health, or investment decisions.<sup>20</sup> As one stark example, [Oster et al. \(2013\)](#) report that only 7 percent of individuals at risk of the hereditary Huntington’s Disease opt for the medical test.

In a laboratory study, [Ganguly and Tasoff \(2017\)](#) give subjects the option to avoid being tested for milder virus infections at a cost of 10 USD. Depending on the type of virus subjects are tested for, 5.2% to 15.6% of individuals are willing to give up 10 USD in order not to be tested, even though the diseases can be treated quite well. [Ganguly and Tasoff \(2017\)](#) also look at demand for non-instrumental information. Giving subjects the option to learn (or avoid) information about potential monetary payments at some cost, they find that some subjects are willing to pay money to avoid learning about the outcome of a lottery whereas others are willing to pay for early resolution of uncertainty. Each of the two groups roughly comprise between 30 and 40 percent of the sample.

[Eil and Rao \(2011\)](#) find evidence for information avoidance when they elicit subjects’ willingness-to-pay of learning their relative rank in terms of IQ and attractiveness. The authors show that subjects who believe to be below average in a category are willing to pay for not learning their true rank in the distribution. Rank is not connected to any direct monetary incentive in that study, but of course, knowing one’s rank may contain instrumental information for future decisions outside the laboratory.

<sup>20</sup> A related phenomenon exists also in the morally relevant domain, compare [Dana et al. \(2007\)](#), [Di Tella et al. \(2015\)](#) or [Falk et al. \(2020\)](#) for empirical evidence.

Our paper also relates to work that studies the demand for non-instrumental information. In [Eliaz and Schotter \(2010\)](#), a majority of subjects choose to pay a fee to obtain information that will only alter their confidence about their decision, but not their decision itself. [Masatlioglu et al. \(2023\)](#) conduct experiments which consider a variety of information structures, including skewed ones, and argue that their results are in line with models of anticipatory utility such as [Caplin and Leahy \(2001\)](#). In [Falk and Zimmermann \(2023\)](#), subjects can choose when to find out about whether they are to receive mild electric shocks. Without anything interesting to do, a large majority chooses to find out immediately, and can thus be classified as “curious”. This is maybe why in our setting, when given the choice, the majority of subjects opt for information, though information avoidance leads to better performance results. More research in this direction would be interesting.

[Ariely et al. \(2009\)](#) demonstrate in experiments conducted in the U.S. and in India, for a variety of different tasks, that high incentives might backfire and reduce performance. For the majority of the tasks administered in their study, performance varies non-monotonically with the compensation offered. Moderate incentives typically deliver the highest performance level. [Dohmen \(2008\)](#) and [Apesteguia and Palacios-Huerta \(2010\)](#) document a similar effect for professional athletes. Building on [Baumeister \(1984\)](#), all these studies stress the phenomenon that individuals may choke under too much pressure induced by high rewards. This renders the fear of choking under pressure that a fraction of our subjects expressed rather plausible.

[Shen et al. \(2015\)](#) document that in certain real effort situations, a small reward that is uncertain and either higher or lower (e.g. 1 USD or 2 USD with equal probability, or a smaller versus a larger amount of candy) may generate better performance results than the fixed higher reward (e.g. 2 USD). In all the settings studied, overall rewards are small. The authors suggest that the uncertainty about the rewards may increase subjects’ excitement, and subsequently, their motivation with which they engage in the task. While this may be true when stakes are low, we consider it unlikely that excitement alone drives our results, in which overall stakes are rather high (about 30 EUR if the piece rate is 1 EUR).

In our ex-post questionnaire, many subjects argue that they decided to avoid information in order to prevent demotivation from a low piece rate realization, while others state that they wanted to avoid feelings of pressure. Excitement from the lottery does not seem to play a major role in the explanations subjects gave. Nevertheless, it would obviously be important to explore in future research the role of different mechanisms for why uncertain incentives can foster performance results. Possibly, the roles of the different mechanisms also depend on context.

## 6. Conclusion

We studied a real effort task in which agents decide whether they wanted to receive instrumental information about their piece rate before starting to work, or not. Our data display that about one third of subjects deliberately decide to forgo information, thus revealing a preference for information avoidance. Furthermore, agents avoiding information achieve considerably better performance outcomes than agents opting for information.

In order to rule out potential causal effects of self-selection on performance, we ran another treatment in which subjects were forced to stay uninformed about their piece rate. Performance again turned out to be significantly higher than under complete information. Information avoidance, even if enforced instead of self-chosen, significantly increases performance.

Two classes of behavioral models, those assuming no belief distortions such as [Bénabou and Tirole \(2002\)](#) or [Carrillo and Mariotti \(2000\)](#) which rationalize information avoidance as a way to overcome self-control problems, and those assuming optimistic belief distortion such as [Brunnermeier and Parker \(2005\)](#), can explain the results up to this point. In order to differentiate between the two, we ran a treatment in which subjects received the expected piece rate with certainty. The data show that subjects in this treatment perform worse than under uncertainty. This result supports the approaches à la [Brunnermeier and Parker \(2005\)](#), under which agents optimistically distort their beliefs away from the objective expectation. In a follow-up treatment, we find indeed that agents perform comparably well as under uncertainty if they receive 0.8 EUR as a secure piece rate. This fits well with the models predicting belief distortions, and sets the findings apart from the predictions generated by models which maintain undistorted beliefs.

Our study documents that giving agents room to design their beliefs may not only be beneficial in contexts such as health (as has been documented by [Oster et al., 2013](#) and [Ganguly and Tasoff, 2017](#)), but also in economic settings of paid effort exertion and performance. While a direct randomization over piece rates may be unpopular with workers and unions, there are other more subtle and perfectly accepted ways of introducing uncertainty about effective pay in a firm. Any type of incentive scheme that introduces interdependencies between workers’ payments generates scope for beneficial belief distortion. Rampant overprovision of effort in contests may have one of its roots in optimistic belief distortion.<sup>21</sup> Additionally, in many, typically large, firms, workers are paid in company shares as part of their salary. Since the individual workers probably have a negligible effect on the evolution of the share’s value, specifically when they work on a lower hierarchy level within the company, this will introduce exogenous uncertainty into the workers’ compensation. Our results suggest that this might be an effective way for a firm to increase workers’ efforts. Employees may distort their beliefs about the firms’ future performance according to their needs, which increases their effort.

The possibility of inducing more effort with less pay is tantalizing. We believe that it offers much scope for further research.

<sup>21</sup> For a detailed survey of the experimental literature on contests, see [Dechenaux et al. \(2015\)](#) who document evidence for overprovision of effort in a large number of different settings.

## Declaration of competing interest

The authors declare the following financial interests/personal relationships which may be considered as potential competing interests: Huck serves as Guest Editor for the Special Issue but has, of course, no editorial involvement in this submission. If there are other authors, they declare that they have no known competing financial interests or personal relationships that could have appeared to influence the work reported in this paper.

## Appendix A

### A.1. Formal derivation of the predictions

We consider an agent who receives a piece rate  $w \in \{w_L, w_H\}$  for working on a task. Completing  $e$  tasks comes at a convex effort cost  $c(e)$ . Hence, in this formulation, effort directly translates into performance, i.e. output. At period 0, the agent faces a lottery where each wage is equally likely to realize and she decides whether to find out her wage or not. In period 1, the agent works on the task and incurs her effort cost. He is paid  $w \cdot e$  in period 2 and does not discount payments. We assume risk neutrality throughout, but emphasize how risk aversion would change the results where appropriate.

#### Information avoidance as a response to self-control problems

The key ingredient of models such as [Bénabou and Tirole \(2002\)](#) or [Carrillo and Mariotti \(2000\)](#) is a temptation problem of the agent. In period 0, for a given wage  $w$ , the optimal effort level  $e^*(w)$  solves

$$w = c'(e^*(w))$$

If, however, the wage is low,  $w = w_L$ , the agent does not exert effort according to  $e^*(w_L)$  in period 1. Instead, the agent gives in to the temptation of “slacking off” and chooses effort level  $\hat{e}(w_L)$  according to:

$$\beta w_L = c'(\hat{e}(w_L))$$

Since  $\beta < 1$  and  $c'(\cdot)$  is increasing by assumption, the agent’s effort at the low wage is lower than ex-ante optimal,  $e^*(w_L) > \hat{e}(w_L)$ .

Now, consider the agent’s decision whether to obtain information about her wage. If she decides to acquire the information, her expected utility will be given by:

$$\frac{1}{2} (w_L \hat{e}(w_L) - c(\hat{e}(w_L))) + \frac{1}{2} (w_H e^*(w_H) - c(e^*(w_H))) \quad (\text{A.1})$$

If instead, she does not obtain information, her expected utility is given by

$$\bar{w} e^*(\bar{w}) - c(e^*(\bar{w}))$$

where  $\bar{w} = \frac{1}{2}(w_L + w_H)$ . For any  $w$ , the agent’s utility  $U(w) \equiv w e^*(w) - c(e^*(w))$  is convex in  $w$ , hence it follows that for  $\beta = 1$ , that is in the absence of a temptation problem, the agent prefers to know her wage because  $\frac{1}{2}U(w_L) + \frac{1}{2}U(w_H) > U(\bar{w})$ . For  $\beta < 1$ , however, the agent’s utility is given by (A.1) which decreases as  $\beta$  decreases. Hence, if  $\beta$  is sufficiently low and  $U(w)$  not too convex, the utility decrease from the loss in motivation is large enough such that:

$$\frac{1}{2} (w_L \hat{e}(w_L) - c(\hat{e}(w_L))) + \frac{1}{2} U(w_H) < U(\bar{w})$$

This implies that the agent prefers to avoid information about her wage. Note, however, that effort when facing the lottery and when facing a sure wage of  $\bar{w}$  is in both cases given by  $e^*(\bar{w})$ . Hence, this model predicts that information avoiders perform no better than agents in our 0.55 EUR (FIXED WAGE) treatment.

#### Information avoidance due to belief distortions

In models which allow for belief distortions such as [Brunnermeier and Parker \(2005\)](#), overall utility is expressed as a weighted average of an anticipatory component and realized (“consumption”) utility. While the anticipatory utility is based on an optimally chosen belief, consumption utility is calculated using the objective belief. Denoting by  $0 \leq \theta < 1$ , the relative weight of the former, the agent’s utility is given by:

$$\theta [w(\pi) e^*(w(\pi)) - c(e^*(\pi))] + (1 - \theta) [\bar{w} e^*(w(\pi)) - c(e^*(w(\pi)))] \quad (\text{A.2})$$

The optimal choice of beliefs has to consider the following trade off: An agent may distort beliefs away from  $\bar{w}$  in order to increase her anticipatory utility. By being more optimistic about her odds to be paid the high wage, she manipulates herself into exerting more effort since effort is determined by the subjective expected wage. However, choosing a belief different from  $\pi = 0.5$  may come at a cost because the agent will exert more effort than what is optimal given  $\bar{w}$ .

We analyze the agent’s problem by starting with the effort choice. Whenever she works under wage uncertainty (whether forced or chosen), she will choose her beliefs  $\pi$ , the likelihood of receiving the high wage. From this, we find  $w(\pi) = \pi w_H + (1 - \pi) w_L$  as the subjective wage. Hence, when the agent works on the task she chooses her effort level taken this distorted wage as given. Importantly, this effort is chosen to maximize the anticipatory utility:

$$w(\pi) = c'(e^*(w(\pi)))$$

When choosing her optimal belief  $\pi^*$ , however, she maximizes the expression in (A.2), that is, she takes into account her anticipatory utility (which only depends on her belief  $w(\pi)$ ) as well as her consumption utility which is calculated based on the objective expected wage  $\bar{w}$ . The first order condition for the belief choice can then be written as:

$$\theta e^*(w(\pi^*)) + (1 - \theta) \frac{de^*(w(\pi^*))}{dw(\pi^*)} (\bar{w} - w(\pi^*)) = 0 \quad (\text{A.3})$$

Due to the convexity of  $c(\cdot)$ ,  $e^*$  is increasing in the wage, hence, for  $\theta = 0$ , we find that  $\bar{w} = w(\pi^*)$  and for  $\theta > 0$  that  $\bar{w} < w(\pi^*)$ . Thus, whenever the agent cares about anticipatory emotions, she distorts her beliefs upwards. We also note that as  $\theta \rightarrow 1$ , the agent will maximally distort beliefs, i.e. select  $\pi^* = 1$ .

What is left to show is that such an agent would find it worthwhile to avoid information. If, on the contrary, she opts for information, she does not distort beliefs and simply chooses effort optimally according to  $w = c'(e(w))$ . We therefore need to have that:

$$\theta U(w(\pi^*)) + (1 - \theta) [\bar{w} e^*(w(\pi)) - c(e^*(w(\pi)))] > \frac{1}{2} U(w_L) + \frac{1}{2} U(w_H)$$

A standard “envelope” argument implies that the L.H.S. is increasing in  $\theta$ . Hence, this condition will eventually be satisfied for a large enough  $\theta$ . This implies that the model can predict information avoidance and at the same time makes a clear prediction regarding the performance level we should observe. Since the agent distorts beliefs her effort is given by  $e^*(w(\pi^*))$  which will be larger than the effort she exerts at a fixed wage of  $\bar{w}$ .

## A.2. A formal model of choking and belief design

We follow the notation in the previous section, with the modification that instead of the assumption that effort  $e(w)$  directly translates into the agent’s output, we write output  $y(w) = e(w)f(w)$ , that is, the product of effort and productivity. The latter can be thought of as the agents’ efficiency in working on the task (i.e. the propensity of making mistakes), potentially depending on the wage she faces. This means that the optimal effort level is chosen as to maximize

$$w(\pi)e f(w(\pi)) - c(e)$$

for a given (potentially distorted) wage  $w(\pi)$ . We thus find the first order condition as:

$$w(\pi)f(w(\pi)) = c'(e^*(w(\pi)))$$

We capture the idea of choking via the productivity  $f(w)$ . Specifically, we assume that for “chokers”,  $f(w)$  is decreasing and concave for wages higher than some threshold wage  $w_c \in (\bar{w}, w_H)$ , and constant otherwise. The decreasing productivity for high wages can be, for example, understood as a higher propensity of making mistakes, or more generally as stress- or pressure-induced adverse effects on performance. Concavity of  $f(w)$  captures the intuition that choking behavior is exacerbated as the wage increases.<sup>22</sup> We furthermore assume that  $\lim_{w \rightarrow w_c} f(w) = 0$ , which then directly implies that if choking is sufficiently severe, there exists a  $w = \hat{w}$  which satisfies  $f'(\hat{w})\hat{w} + f(\hat{w}) = 0$ . If  $w_c < \hat{w} < w_H$ , this captures the essence of the agent’s choking behavior for sufficiently high wages.  $\hat{w}$  is the wage which maximizes the agents’ effort. Furthermore, it follows that effort is increasing in  $w$  for wages lower than  $\hat{w}$ , and decreasing for wages above  $\hat{w}$ .

As before, the agent now maximizes the weighted sum of anticipatory utility and consumption utility. In our setting, the belief choice directly corresponds to a choice of the subjective wage. To ease exposition, we therefore suppress the dependence of  $w$  on  $\pi$  and allow the agent to choose  $w^* \in [w_L, w_H]$  directly. The agent maximizes:

$$\theta [w e^*(w) f(w) - c(e^*(w))] + (1 - \theta) [\bar{w} e^*(w) f(w) - c(e^*(w))]$$

We then obtain the first order condition as:

$$\frac{f'(w)}{f(w)} w + \theta - \left( \frac{f'(w)}{f(w)} + \frac{e^{*'}(w)}{e^*(w)} \right) (1 - \theta)(w - \bar{w}) = 0$$

From this expression we notice the following: If there is no choking, i.e.  $f(w)$  is constant for all  $w \leq w_H$ , the condition collapses to the first order condition in (A.3) above.

The more interesting case is the one with choking. Since  $\frac{e^{*'}(w)}{e^*(w)} > 0$  for  $w \leq \hat{w}$ , in order for there to exist a solution  $w^* > w_c$ , it must be the case that there exists a wage  $x \in (w_c, \hat{w})$  which satisfies

$$\theta \geq - \frac{f'(x)\bar{w}}{f(x) + f'(x)(x - \bar{w})}$$

If there exists no such wage, the optimal wage  $w^*$  lies in the interval  $[\bar{w}, w_c]$  and solves (A.3). If  $\theta$  is large enough the condition will be satisfied. Indeed, for  $\theta \rightarrow 1$ , the condition can be rewritten as

$$f(x) + f'(x)x \geq 0$$

which is true for any  $x \leq \hat{w}$ , due to the concavity of  $f(x)$ . Hence, for agents who choke under pressure, we also predict belief distortions. However, these agents will never distort beliefs above  $\hat{w}$ , because for any  $w > \hat{w}$  there is a  $w < \hat{w}$  which induces the

<sup>22</sup> The assumption that there is no choking below the average wage  $\bar{w}$  allows us to focus attention on upward belief distortions which we see in the data.

**Table A.1**  
Summary statistics of individual characteristics.

|               | All   | FULL INFO | INFO CHOICE | NO INFO | $w = 0.55$ | $w = 0.80$ |
|---------------|-------|-----------|-------------|---------|------------|------------|
| % Female      | 0.326 | 0.313     | 0.326       | 0.313   | 0.362      | 0.319      |
| % Non-German  | 0.116 | 0.042     | 0.105       | 0.125   | 0.191      | 0.128      |
| Age           |       |           |             |         |            |            |
| mean          | 23.18 | 22.17     | 23.2        | 22.42   | 24.20      | 23.98      |
| median        | 23    | 21.5      | 23          | 22      | 23.5       | 23         |
| # of semester |       |           |             |         |            |            |
| mean          | 4.45  | 3.60      | 4.96        | 3.65    | 4.35       | 5.21       |
| median        | 3     | 3         | 4           | 3       | 4          | 4          |

Note: In the category age, we exclude one subject (treatment  $w = 0.55$ ) who reports an (implausible) age of 97. In the category semester, we exclude two subjects (one in treatment  $w = 0.55$  and one in treatment INFO CHOICE) who did not report any value.

**Table A.2**  
Summary statistics of individual characteristics in treatment INFO CHOICE, depending on whether subject avoided information or not.

|               | Info avoiders | Info receivers | $p$ -value <sup>a</sup> |
|---------------|---------------|----------------|-------------------------|
| % Female      | 0.3           | 0.339          | 0.816                   |
| % Non-German  | 0.067         | 0.123          | 0.496                   |
| Age           |               |                |                         |
| mean          | 22.8          | 23.385         |                         |
| median        | 22            | 23             | 0.2379                  |
| # of semester |               |                |                         |
| mean          | 4.724         | 5.062          |                         |
| median        | 3             | 5              | 0.2406                  |

Note: In the category semester, we exclude one subject (Info Avoider) who did not report any value.

<sup>a</sup> For the first two rows, the  $p$ -value stems from a Fisher-exact test, for the last two rows, a Wilcoxon rank-sum test is used.

same effort at a higher productivity.

In order to make statements about the agent's output, i.e. the number of correctly solved tasks, we need to consider  $y(w^*)$  which we defined as  $y(w^*) = e^*(w^*)f(w^*)$ . We first note that since  $e^*(w)$  changes sign at most once (effort increases in the wage for low wages and decreases for wages above  $\hat{w}$ ), for chokers  $y(w)$  is concave with a unique maximum at  $z \in (w_c, \hat{w})$ . This implies that  $y(\hat{w}) > y(w_H)$ . Hence, for optimal wages  $w^* \in [z, \hat{w}]$ , the model predicts that performance is higher at an optimally distorted wage  $w^*$  than at the certain wage of  $w_H$ . Naturally, this will even be the case for some wages below  $z$ , depending on how large the negative effects of choking at  $w = w_H$  are.<sup>23</sup>

Finally, we can use this result to argue why, in a heterogeneous population where only some agents choke at high wages, coarse information structures can do better than any fixed wage  $w \in [w_L, w_H]$ . For simplicity consider the case where  $\theta$  is close to 1. With wage uncertainty, non-choking agents optimally distort their beliefs such that  $w(\pi) \approx w_H$ , whereas choking agents choose beliefs such that  $w(\pi) \approx \hat{w}$ . Hence, while both types of agents distort their beliefs differently, the aggregate output of a group of agents consisting of both types will be higher than it could be under a fixed wage of either  $w_H$  or  $\hat{w}$ , or indeed any other fixed wage.

### A.3. Additional tables and robustness checks

See Tables A.1–A.4.

## Appendix B. Supplementary materials

Supplementary material related to this article can be found online at <https://doi.org/10.1016/j.eurocorev.2025.104965>.

<sup>23</sup> By using more specific functional forms for  $f(w)$  and  $c(e)$ , one could derive analytical expressions for the optimal wage and how it depends on  $\theta$ . However, the main point of this section is to demonstrate the possibility of higher performance under uncertainty, hence few new insights would be gained from such an exercise.



**Table A.3**  
OLS regressions of number of correctly solved tasks, checking for selection effects.

|                                             | (1)                 | (2)                 | (3)                 |
|---------------------------------------------|---------------------|---------------------|---------------------|
| $w = 1$                                     | 7.848***<br>(2.361) | 7.862***<br>(2.423) | 7.808***<br>(2.421) |
| $w$ uncertain                               | 12.31***<br>(2.361) | 12.39***<br>(2.563) | 12.74***<br>(2.401) |
| $w = 0.1 \times \text{FULL INFO}$           | 2.981<br>(2.515)    | 2.937<br>(2.563)    | 2.879<br>(2.561)    |
| $w = 1 \times \text{FULL INFO}$             | 0.675<br>(2.599)    | 0.641<br>(2.647)    | 0.626<br>(2.644)    |
| $w$ uncertain $\times$ No INFO              | -1.979<br>(2.208)   | -2.044<br>(2.221)   | -2.449<br>(2.225)   |
| Age                                         |                     | 0.151<br>(0.279)    | 0.176<br>(0.279)    |
| Female = 1                                  |                     | 2.894<br>(1.501)    | 3.096**<br>(1.509)  |
| Non-German = 1                              |                     | -0.925<br>(2.422)   | -1.052<br>(2.357)   |
| Semester                                    |                     | -0.152<br>(0.288)   | -0.196<br>(0.290)   |
| Constant                                    | 17.69***<br>(1.604) | 14.06***<br>(5.939) | 13.66***<br>(5.943) |
| Observations                                | 191                 | 191                 | 190                 |
| R-squared                                   | 0.176               | 0.194               | 0.199               |
| $H_o(\{w = 1\} = \{w = \text{uncertain}\})$ | $p = 0.0699$        | $p = 0.0688$        | $p = 0.0497$        |
| $H_o(\text{No Selection})^a$                | $p = 0.5187$        | $p = 0.5248$        | $p = 0.4688$        |

Note: The table reports coefficients from an OLS regression of the number of correctly solved tasks for the treatments FULL INFO, INFO CHOICE, No INFO on dummies for the different wages and interactions with treatment dummies. Column (1) includes no additional control, whereas column (2) and (3) include the additional controls, as described in Table A.1. In column (3) we exclude one subject who does not report the number of semesters studied. In column (2) we include this subject by setting the number of semester studied to zero. <sup>a</sup>: To test for no selection effects, we test the joint hypothesis that the coefficients of all three interaction terms are jointly equal to zero. \*\*\*  $p < 0.01$ , \*\*  $p < 0.05$ .

**Table A.4**  
OLS regressions of number of correctly solved tasks.

|                                               | (1)                  | (2)                  | (3)                  |
|-----------------------------------------------|----------------------|----------------------|----------------------|
| $w = 0.1$                                     | -5.761***<br>(1.827) | -5.400***<br>(1.870) | -5.562***<br>(1.891) |
| $w = 0.8$                                     | 3.213<br>(1.928)     | 3.582<br>(1.945)     | 3.412<br>(1.966)     |
| $w = 1$                                       | 1.174<br>(1.864)     | 1.186<br>(1.888)     | 1.024<br>(1.911)     |
| $w$ uncertain                                 | 4.122**<br>(1.726)   | 4.215**<br>(1.757)   | 4.162**<br>(1.786)   |
| Age                                           |                      | 0.0930<br>(0.193)    | 0.0891<br>(0.194)    |
| Female = 1                                    |                      | 0.936<br>(1.196)     | 0.949<br>(1.206)     |
| Non-German = 1                                |                      | -0.223<br>(1.771)    | 0.108<br>(1.800)     |
| Semester                                      |                      | -0.341<br>(0.211)    | -0.340<br>(0.213)    |
| Constant                                      | 24.66***<br>(1.363)  | 23.57***<br>(4.448)  | 23.78***<br>(4.474)  |
| Observations                                  | 285                  | 285                  | 282                  |
| R-squared                                     | 0.132                | 0.143                | 0.145                |
| $H_o(\{w = 0.8\} = \{w = 1\})$                | $p = 0.2751$         | $p = 0.2044$         | $p = 0.2064$         |
| $H_o(\{w = 0.8\} = \{w = \text{uncertain}\})$ | $p = 0.5985$         | $p = 0.7172$         | $p = 0.6683$         |
| $H_o(\{w = \text{uncertain}\} = \{w = 1\})$   | $p = 0.0758$         | $p = 0.0687$         | $p = 0.0603$         |

Note: The table reports coefficients from an OLS regression of the number of correctly solved tasks on dummies for the different wages, using the fixed wage of  $w = 0.55$  as the omitted category. Column (1) includes no additional control, whereas column (2) and (3) include the additional controls, as described in Table A.1. In column (3) we exclude three subjects, two of which do not report the number of semesters studied, an one reports an (implausible) age of 97, whereas in column (2) we include them by setting the number of semester studied to zero and the age to the median age of 23 in our sample. \*\*\*  $p < 0.01$ , \*\*  $p < 0.05$ . Tests of coefficients of ( $w = 0.1$ ) against all other wages always significant at  $p < 0.01$ .

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